## MLDS: Maximum Likelihood Difference Scaling in R

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## Psychophysics,

qu'est-ce que c'est ?
A body of techniques and analytic methods to study the relation between physical stimuli and the organism's (classification) behavior to infer internal states of the organism or their organization.

Gustav Fechner (I80I-I887)


Difference scaling is a psychophysical procedure used to estimate a perceptual scale for stimuli distributed along a physical continuum.

Example: series along a line in tristimulus space, What are the perceptual distances between samples?


Maloney \& Yang, 2003

## Example: VQ compressed images,

Up to what compression rate can the observer detect no loss of image quality?


15:1


18:1


30:1


Charrier, Maloney, Cherifi \& Knoblauch, submitted

## Example: Correlation in scatterplots



## Difference Scaling: Experimental Procedure

From a set of p stimuli, $\left\{I_{1}<I_{2}<\ldots<I_{p}\right\}$,
a random quadruple, $\left\{I_{a}, I_{b} ; I_{c}, I_{d}\right\}$, is chosen (w/out replacement) and presented to the observer as in this example,


Between which pair (upper/lower) is the perceived difference greatest?


For a sequence of $\mathbf{p}$ stimuli, there are $\binom{p}{4}$ non-ovelapping quadruples:

210 quadruples for $p=10$
330 quadruples for $p=11$, etc.

At the termination of an experiment (about 15 minutes), the data are stored in a 5 column data frame. 6 lines of an example are shown here.

One column for the response indicating whether the upper (I) or lower (0) pair was chosen.
4 columns for the indices of the stimuli in the sequence, $1: p$

The aim of the Maximum Likelihood Difference Scaling (MLDS) procedure is to estimate scale values, $\left(\psi_{1}, \psi_{2}, \ldots, \psi_{p}\right)$, that best capture the observer's judgments of the perceptual difference between the stimuli in each pair.

The MLDS package, available on CRAN, provides tools for performing this analysis in R. An example scale obtained from an observer for the "apples" sequence of VQ compressed images is shown on the right:


## The decision model

Given a quadruple, $\mathbf{q}=(a, b ; c, d)$, from a single trial, we assume that the observer chooses the upper pair to be further apart when
$\Delta(a, b ; c, d)=\left|\psi_{d}-\psi_{c}\right|-\left|\psi_{b}-\psi_{a}\right|+\epsilon>0$,
where $\psi_{i}$ are estimated scale values, $\epsilon \sim N(0, \sigma)$, and $\sigma$ a scale factor.

## Estimation of Scale Values

Maloney and Yang (2003) used a direct method for estimating the maximum likelihood scale values,

$$
L(\mathbf{\Psi}, \sigma)=\prod_{k=1}^{n} \Phi\left(\frac{\delta\left(\mathbf{q}^{k}\right)}{\sigma}\right)^{1-R_{k}}\left(1-\Phi\left(\frac{\delta\left(\mathbf{q}^{k}\right)}{\sigma}\right)\right)^{R_{k}}
$$

where

$$
\begin{aligned}
& \Psi=\left(\psi_{2}, \psi_{3}, \ldots, \psi_{p-1}\right) \\
& \delta\left(\mathbf{q}^{k}\right)=\left|\psi_{d}-\psi_{c}\right|-\left|\psi_{b}-\psi_{a}\right|
\end{aligned}
$$

$\Phi$ is the cumulative standard Gaussian (a probit analysis)
$R_{k}$ is $0 / I$ if the judgment is lower/upper
$\psi_{1}=0, \psi_{p}=1$ for identifiability,
leaving $p-1$ parameters to estimate

## Estimation of Scale Values

The problem can also be conceptualized as a GLM.
Each level of the stimulus is treated as a covariate in the model matrix, taking on values of 0 or $\pm 1$ in the design matrix, depending on the presence of the stimulus in a trial and its weight in the decision variable, with absolute value signs removed.

|  |  |  |  |  |  |
| ---: | ---: | ---: | ---: | ---: | ---: |
|  | resp | S1 | S2 | S3 | S4 |
| 1 | 0 | 4 | 8 | 2 | 3 |
| 2 | 1 | 2 | 3 | 6 | 11 |
| 3 | 1 | 2 | 6 | 7 | 10 |
| 4 | 0 | 4 | 11 | 1 | 2 |
| 5 | 0 | 9 | 11 | 7 | 8 |
| 6 | 0 | 7 | 10 | 1 | 3 |\(\quad\left(\begin{array}{rrrrlllllll}p_{1} \& p_{2} \& p_{3} \& p_{4} \& p_{5} \& p_{6} \& p_{7} \& p_{8} \& p_{9} \& p_{10} \& p_{11} <br>

0 \& 1 \& -1 \& -1 \& 0 \& 0 \& 0 \& 1 \& 0 \& 0 \& 0 <br>
0 \& 1 \& -1 \& 0 \& 0 \& -1 \& 0 \& 0 \& 0 \& 0 \& 1 <br>
0 \& 1 \& 0 \& 0 \& 0 \& -1 \& -1 \& 0 \& 0 \& 1 \& 0 <br>
1 \& -1 \& 0 \& -1 \& 0 \& 0 \& 0 \& 0 \& 0 \& 0 \& 1 <br>
0 \& 0 \& 0 \& 0 \& 0 \& 0 \& 1 \& -1 \& -1 \& 0 \& 1 <br>
1 \& 0 \& -1 \& 0 \& 0 \& 0 \& -1 \& 0 \& 0 \& 1 \& 0\end{array}\right)\)

For model identifiability, we drop the first column (fixing $\psi_{1}=0$ and $\sigma=1$ ).

## Estimation of Scale Values

$>$ kk.ix <- make.ix.mat(kk)
$>$ head(kk.ix)
resp stim. 2 stim. 3 stim. .4 stim. 5 stim. 6 stim. 7 stim. 8 stim. 9 stim. 10 stim. 1 I

| I | I | 1 | 0 | -I | 0 | -I | 0 | 1 | 0 | 0 | 0 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2 | I | 0 | 0 | -I | 0 | -I | 0 | 0 | I | 0 | 0 |
| 3 | I | I | -1 | 0 | 0 | 0 | -I | 0 | I | 0 | 0 |
| 4 | 1 | 1 | 0 | 0 | -1 | -1 | I | 0 | 0 | 0 | 0 |
| 5 | 0 | 0 | -1 | 0 | 0 | -1 | 1 | 0 | 0 | 0 | 0 |
| 6 | 0 | 0 | 0 | 0 | -I | -1 | 0 | 0 | I | 0 | 0 |

$$
\eta(\mathrm{E}[Y])=X \beta
$$

> glm(resp ~ . - I, family = binomial( "probit" ), data = kk.ix)

## The MLDS package

The MLDS package provides a modeling function, mlds() , that is essentially a wrapper for either glm() or optim(), and will enable estimation of the perceptual scale values, given a data frame with the previously described structure.

```
mlds(data, stimulus, method = "glm", lnk = "probit",
    opt.meth = "BFGS", opt.init = NULL,
    control = glm.control(maxit = 50000, epsilon = 1e-14),
    ... )
```

It outputs an S3 object of class 'mlds' which can be examined further using several method functions:
summary, plot, predict, fitted, logLik, AIC and boot

As a running example，we consider data sets from an experiment in which one observer judged differences in correlation between scatterplots for II levels of correlation：

$$
c(\operatorname{seq}(0,0.9, \text { len }=10), 0.98)
$$

100 points generated with：
MASS：：：mvrnorm（I00，mu＝c（0，0），Sigma $=$ matrix（c（r，0，0，r），2，2））

| $r=0$ | $r=0.1$ | $r=0.2$ | $r=0.3$ |
| :---: | :---: | :---: | :---: |
| : | - | : |  |
| $r=0.4$ | $r=0.5$ | $r=0.6$ | $r=0.7$ |
|  |  | aciso | 荡 |
| $r=0.8$ | $r=0.9$ | $r=0.98$ |  |
| 里密 | $: 94 \times 0$ |  |  |
|  |  |  | oblauch |

The first author ran himself on 330 trials on 3 separate days, generating data sets kkI, kk2, kk3, available in the package. A typical trial is indicated below.

Between which pair, lower/upper, is the difference greatest?

```
> library(MLDS)
> data(kkl)
> data(kk2)
> data(kk3)

The data sets have class 'mlds.df' that inherits from 'data.frame'. It differs in including two attributes, "stimulus" and "invord".
```

> str(kkl)
Classes 'mlds.df' and 'data.frame': 330 obs. of 5 variables:
\$ resp:int l 0 0 0 | | | | 0 | ...
\$ SI :int 267666| 323 ...
\$ S2 :int 4997792554 ...
\$ S3 :int 6 | 2 2 | | 8 IO 7 5 ...
\$ S4 :int 8435359 || 8 IO ...

- attr(*,"invord")= logi FALSE TRUE TRUE TRUE TRUE
TRUE .
    - attr(*,"stimulus")= num 0.0 0.I 0.2 0.3 0.4 ...
stimulus is a numeric vector of the physical stimulus levels invord is a logical vector indicating whether on each trial the higher scale values were on the botton or top.

```

It is sometimes necessary to reorder the pairs so that the higher physical values are after the lower ones.

This is conveniently done with the SwapOrder() function, which uses the "invord" attribute, if present.

The Rbind() function combines data sets and their attributes.
> kk <- SwapOrder( Rbind( kkl, kk2, kk3 ) )
to produce one large object.
\(>(\) kk.mlds <- mlds(kk) \()\)

\section*{Perceptual Scale:}
\(\begin{array}{lllllllll}0 & 0.1 & 0.2 & 0.3 & 0.4 & 0.5 & 0.6 & 0.7 & 0.8\end{array}\)
 \(0.9 \quad 0.98\)
4.42975 .5739
sigma:
[I] I
\(>(\) kk.mlds2 <- mlds(kk, method \(=\) "optim", opt.init \(=c(s e q(0, I\), len \(=I I), 0.2)))\)
Perceptual Scale:

sigma:
[I] 0.175
> plot(kk.mlds, standard.scale =TRUE,
cex = I.7, pch = I6, col = "black")
> lines(kk.mlds2, lwd = 2)


\section*{Bootstrap Errors on mlds Scale Values}

We have included a function for estimating bootstrap standard errors on the scale values
boot.mlds(x, nsim, ...)
where \(x\) is of class 'mlds' and nsim is the number of bootstrap repetitions.

The fitted probabilities are used with rbinom() to generate new responses which are fitted using mlds() to generate new bootstrap scale values.

These boostrap scale values are then used to generate the error bars (10 000 replications)


\section*{Scale validation}
I. Ordering property: Observer must be able to reliably order stimuli, \(\left(I_{1}, I_{2}, \ldots, I_{p}\right)\), in agreement with scale values, \(\left(\psi_{1}, \psi_{2}, \ldots, \psi_{p}\right)\). (usually not formally tested and evidently satisfied).
2. Six-point property: Given any two groups of three intervals, \((a, b, c)\) and \(\left(a^{\prime}, b^{\prime}, c^{\prime}\right)\),

if \(a b \succ a^{\prime} b^{\prime}\) and \(b c \succ b^{\prime} c^{\prime}\) then \(a c \succ a^{\prime} c^{\prime}\).
Maloney and Yang (2003) proposed a resampling method for testing the six-point property that we have implemented in MLDS.

D .H. Krantz, R. D. Luce, P. Suppes., and A. Tversky. Foundations of Measurement (Vol. 1): Additive and Polynomial Representation. Academic Press, New York, 1971.

\section*{Scale validation: Six-point test}

We have written a function to estimate the likelihood of the six-point choices from an 'mlds' object and then using the fitted probabilities to generate new responses to be fit, that permit us to compare the observed likelihood with those based on a large number of resamplings.
kk.6pt <- simu.6pt(kk.mlds, nsim = I0000)
returns the six-point likelihood for the observed data as well as the 10000 resampled likelihoods.
\[
\begin{aligned}
& L_{6 p t}=-425 \\
& p=0.85
\end{aligned}
\]


\section*{Future Directions}
I. Formula interface:

The current fitting procedure requires the estimation of \(p-1\) parameters. The fitted scale may suggest a simpler parametric form with fewer parameters.
kk.fun <- mlds.function( \(\sim s x^{\wedge} p, p=c(4,0.2)\), data \(\left.=k k\right)\)
takes a one-sided formula with parameters, p and sigma, yielding only two parameters.
plot(kk.mlds, standard.scale \(=\) TRUE,
cex = I.7, pch = I6, col = "black")
lines(kk.fun\$stimulus, kk.fun\$pscale, lwd = 2)
> -2 * kk.fun\$logLik + 2 * 2
[I] 656.8796
> AIC(kk.mlds)
[I] 632.7912


\section*{Future Directions}

Another example of the formula interface:
> kk.fun2 <- mlds.function(~p[I] * (sx + abs(sx - p[2])) - p[I] *p[2],
\[
p=c(0.9,0.3,0.2), \text { data }=k k)
\]
> kk.fun2\$par
[I] 0.5I36073 0.3490738
> kk.fun2\$sigma
[I] 0.1391141
> AIC(kk.mlds)
[I] 632.7912
> -2 * kk.fun2\$logLik + 3 *2
[I] 658.6734


\section*{Future Directions}
2. Mixed effects models:

We may want to introduce random effects to account for differences in sensitivity between runs or observers.

We have been experimenting with the Ime4 package for this. For example,
```

library(lme4)
Run <- factor(rep(paste("R", 1:3, sep = ""), each = 330))
kk.ix <- make.ix.mat(kk) \#generate the data frame for glm
kk.lmer <- lmer(resp ~ . + (1 | Run) - 1, data = kk.ix,
family = binomial("probit"))

```
> summary(kk.Imer)
Generalized linear mixed model fit using Laplace
Formula: resp ~ . + (I | Run) - I
Data: kk.ix
Family: binomial(probit link)
AIC BIC logLik deviance
> AIC(kk.mlds)
630.8 684.7-304.4 608.8
[I] 632.7912
Random effects:
Groups Name Variance Std.Dev.
Run (Intercept) 0.130260 .36092
number of obs: 990, groups: Run, 3
Estimated scale (compare to I) 28.98914
Fixed effects:
Estimate Std. Error z value \(\operatorname{Pr}(>|z|)\)
\begin{tabular}{llllll} 
stim. 2 & 0.05725 & 0.14189 & 0.403 & 0.687
\end{tabular}
\(\begin{array}{llllll}\text { stim. } 3 & 0.23181 & 0.15474 & 1.498 & 0.134\end{array}\)
\(\begin{array}{llllll}\text { stim. } 4 & 0.13198 & 0.16748 & 0.788 & 0.431\end{array}\)
\(\begin{array}{lllll}\text { stim. } 5 & 0.84308 & 0.18613 & 4.529 & 5.91 \mathrm{e}-06\end{array} * * *\)
stim. 6 I.715I6 \(0.21032 \quad 8.1553 .49 \mathrm{e}-16\) ***
stim. \(7 \quad 2.34077 \quad 0.23529 \quad 9.948<2 \mathrm{e}-16 * * *\)
stim. \(8 \quad 2.87728 \quad 0.2650 \mathrm{I} \quad 10.857<2 \mathrm{e}-16\) ***
stim. \(9 \quad 3.671650 .3067 \mathrm{I}\) II.97I \(<2 \mathrm{e}-16\) ***
stim.10 \(4.43316 \quad 0.35794\) 12.385 < 2e-16 ***
stim.II \(5.39385 \quad 0.43525\) I2.393 < 2e-I6 ***
```

